Predicting drowning from sea and weather forecasts: development and validation of a model on surf beaches of southwestern France

Éric Tellier, Bruno Simonnet, Cédric Gil-Jardiné, Marion Lerouge-Bailhache, Bruno Castelle, Rachid Salmi

INTRODUCTION

According to the 2017 Global Burden of Disease study, drowning is a major cause of non-intentional deaths from injury worldwide. In France, the national public health agency (Santé Publique France) performs a national study every 3 years, registering all cases of drowning leading to hospitalisation or death between 1 June and 15 September of the studied year. In 2015, this study reported 1266 drownings, with 637 (50.3%) occurring along the seashore. In a previous study on the surf beaches of Gironde, southwestern France, 576 people required rescue over 6 years; there were 24 fatalities due to drowning. In terms of the length of the coastline, the annual mean was 3.3 deaths/100 km, a rate comparable to the highest recorded along the US coastline.

The Gironde coast is a 126-km-long stretch of sandy beaches (figure 1) exposed to high-energy waves that drive intense, narrow seaward-flowing rip currents. Preventative measures taken to reduce rip current from mid-June to mid-September. On weekends in May and June, some areas are watched, depending on local authorities. The mayor is responsible for beach supervision, which is regionally coordinated by the departmental prefect in collaboration with the prehospital care department of Bordeaux University Hospital. During high season, rescue helicopters are on standby (figure 1). On low-season weekends, one helicopter may be on duty, depending on the regional authority.

Models predicting the coast-wide life-risk of drowning would be useful if they enhanced the preventative measures taken to reduce risk. Predictive models of rip currents have been implemented in Florida, Puerto Rico, Mexico, India and Great Britain. These models were based on physics, modelling the occurrence and the speed of rip currents. They have been evaluated both retrospectively and in the field using hindcasts. To the best of our knowledge, they have not been prospectively evaluated using forecasts, and model predictions have not been compared with actual drownings.

The number of people exposed to a rip hazard is directly related to the number of swimmers and other water users, and therefore linked to beach attendance. Attendance rises on holidays, weekends, and with increased air temperature and less cloud cover; the number of bathers reflects air and water temperatures and (possibly) wind speed. As the risk of drowning is a combination of the hazard per se and exposure to it, and as the latter is poorly quantified, we created a model including parameters reflecting exposure to rip currents. We assessed

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whether drownings off Gironde beaches could be anticipated using a coast-wide risk prediction model based on forecast metocean conditions.

MATERIALS AND METHODS

Study setting

We performed an observational study along the French Atlantic coastline of Gironde (figure 1). The coast is meso-macrotidal, with spring tide range reaching 5 m. Summer-averaged significant wave height and peak period are approximately 1.3 m and 9 s, respectively.22 Deep rip channels incise the inner tidal bar, with an average spacing of c. 400 m, through which intense rip currents can flow. In typical summer wave conditions, rip current activity is maximised for long period waves and shore-normal incidence between low-tide and mid-tide.23 Even for waves of approximately 1 m, mean rip current speed can reach 1 m/s.23

We first developed a model based on medical emergency calls from beaches, along with observed metocean conditions, in 2011, 2012 and 2013. We evaluated only the bathing season (April–October). We tested the model to assess whether it accurately predicted events that occurred from April to October in 2015, 2016 and 2017, using metocean forecasts. We used the RiGoR guidelines24 to address common sources of bias in risk-prediction models, and we adhered to the Strengthening the Reporting of Observational Studies in Epidemiology statement for observational studies.25

Data sources

Medical emergency calls

In Gironde, medical emergency calls either from a bystander or a lifeguard are received by a single medical emergency call centre (Service d’Aide Médicale d’Urgence). During each call, a physician records all information given by the caller, paramedics and (when applicable) prehospital care teams. All calls dealing with rescue from water or drowning were included in the data for this study; these were the events of interest. ‘Rescue’ refers to a need for evacuation from the water,14 and ‘drowning’ refers to respiratory impairment caused by submersion or immersion, as defined by the WHO.26 Both events were considered as adverse water events. We excluded calls lacking victims, training calls and duplicates. As every instance of a need for medical advice or a prehospital care team triggered a call, we considered that all events of importance would be identified. Information on every call was carefully read to avoid errors. Intentional drownings and drownings associated with known diseases (eg, seizure) were excluded.

Environmental conditions

Hourly tidal data were modelled by the ‘Service Hydrographique et Océanographique de la Marine’ (SHOM, authorisation no. 296/2014) using the Lacanau shore as the reference. Lacanau is located in the approximate centre of the study area; according to the SHOM, the maximum tide phase lag over the entire study area is approximately 15 min. Wave conditions were measured every 30 min by the Centre d’Archivage National de Données de Houle In-Situ buoy27 located at 04°43.150′N and 001°26.800′W (figure 1). The wave propagation time from the buoy to the coast is about 1 hour. Observed and forecast meteorological and wave conditions were provided by Météo-France, the French national meteorological service. We used data from the Cap Ferret weather station; Météo-France claims that these well-represent the weather along the entire Gironde coast. Retrospectively, forecasts were not available and observed data were used for 2011–2013. Forecast data, collected prospectively, were available for up to 3 days and at 3 hour steps (7:00 am UTC ±00:00, 10:00 am, etc.). Weather and sea forecasts were Météo-France expert data based on the AROME and WW3 models, respectively. We recorded sea height, the wave height, period and direction. We also recorded wind speed and direction, air and water temperatures, and cloud cover. Other factors influencing beach attendance were the season and type of day. High season was defined as the period from 15 June to 15 September, when most lifeguard stations are open. We distinguished between weekdays, weekends and holidays.

Statistical methods

We fitted two logistic regression models: a ‘daily model’ predicting the overall coast-wide risk of at least one adverse water event on a given day, and a ‘3-h-step model’ predicting the risks at different times of the day (9:30 am–12:29 pm, 12:30–3:29 pm, 3:30–6:29 pm and 6:30–9:30 pm; all local times). Given above mentioned differences in the environmental data collection modes between the training and validation periods, we checked data consistency both visually and using the Wilcoxon-Mann-Whitney and Student’s t-tests.

Days for which metocean data were lacking were removed from the analysis. Prospective cohort data (including variable selection) were not used during model development. We transformed the wave parameters: the wave factor ($W_F$) is the product of significant wave height and peak wave period, and the wave incidence factor $D_F$ is defined by the equation (1), with $\theta$ the mean wave direction (in degrees):

$$D_F = \cos^4 (\theta - 278) \tag{1}$$

278° is the mean angle of the normal to the coastline of the studied area. Therefore, $D_F$ range is 0–1 and is maximal when the wave direction is perpendicular to the coast. This transformation emphasizes on small variations around the shore normal and made $D_F$ log-linear. We categorised non-log-linear quantitative variables (temperatures, wave factor and sea height); these were first divided into quantiles and then reduced using
the Akaia Information Criterion (AIC) in a multivariate context. Model selection used the AIC to perform interaction checks; we tested all possible models. ORs with 95% CIs were computed as bootstrap estimates. We checked that residual autocorrelation was absent. Goodness of fit was assessed using the Le Cessie-van Houwelingen test. Calibration was assessed graphically employing a locally weighted, least-squares smoother and the Spiegelhalter Z-test. Discriminatory power of a daily drowning occurrence according to the final risk scale using the quintiles of the fitted probabilities. All analyses employed R software running the RMS31 and pROC packages.

**RESULTS**

Retrospective data were lacking for 77 days because of a buoy failure, and for 26 prospective days (21 because of data-link loss and 5 because of server unavailability). We analysed 563 days during 2011–2013; 242 adverse water events (136 rescues and 106 drownings) were reported on 108 different days. In 2015–2017, data were available for 612 days; there were 232 adverse water events (155 rescues and 77 drownings) on 104 different days (Table 1). Demographics were similar between the two periods, with a sex ratio of 0.61 F/M, and a median age of 23 (quartiles (18–42)). All retrospective and prospective cohort data were consistent, except for wind speed, which differed between the prospective and retrospective periods. Both were excluded from prospective analyses.

The final, predictive, daily coast-wide life-risk model included wave and wave incidence factors, air temperature, type of day and season (Table 2). Water temperature and tidal range were not retained in the model. The model predicting risk at 3-hour steps featured sea height, wave parameters, air temperature, time of day, type of day and season (see online supplemental appendix table A1). Variation in the daily model was attributable principally to air temperature (proportion of the overall χ² value, 40.9%), wave factors (21.7%), and time of day (16.2%) (figure 2). The principal 3-hour-step model predictors were air temperature (28.9%), the time of day (17.8%), and wave factors (12.6%) (online supplemental appendix figure A1). The probability of a daily drowning occurrence according to the final model is given by the equation (2), with D, the wave incidence factor, W, the wave factor and Tair the air temperature.

\[
 \hat{\gamma} = -7.83 + 1.65[Tair,c] + 2.91[Tair > 23.5] + 1.14[\text{highseason}] - 1.45[\text{weekend}] + 0.363[\text{weekend}] + 3.12 \times D + 1.86[W, \leq 9.2] + 1.96[W, > 9.2]
\]

The daily model had areas under the curves (AUCs) of 0.88 (95% CI 0.84 to 0.91) for 2011–2013 and 0.82 (95% CI 0.78 to 0.86) for 2015–2017 (figure 3). The 3-hour risk model had AUCs of 0.89 (95% CI 0.87 to 0.92) for 2011–2013 and 0.85 (95% CI 0.81 to 0.88) for 2015–2017 (figure 3). Model outcomes did not differ when forecasts for 1, 2 and 3 days were used (p>0.05). Both models were well calibrated in terms of retrospective data (goodness-of-fit test p=0.20 for the daily model, p=0.53 for the 3-hour-step model). Both models exhibited significant p values on Spiegelhalter Z-testing of prospective data, evidencing a lack of calibration: the daily model tended to overpredict days with risks of drowning >0.5; the 3-hour step model overpredicted risks as low as 0.1.

Using prospective data with 3-day forecasts, we found that assessment of the coast-wide risk of water adverse event using the five-level scale missed 1 of 158 days featuring a rescue at

### Table 1 Description of days without and with adverse water events

<table>
<thead>
<tr>
<th></th>
<th>2011–2013 Days without events* (n=455)</th>
<th>2015–2017 Days without events (n=513)</th>
<th>2015–2017 Days with events (n=104)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Wave factor, m/s</td>
<td>11.3 (5.6–14.6)</td>
<td>10.3 (6.5–12.9)</td>
<td>11.3 (5.4–14.4)</td>
</tr>
<tr>
<td>Wave incidence factor</td>
<td>0.80 (0.69–0.99)</td>
<td>0.89 (0.85–0.99)</td>
<td>0.80 (0.74–0.97)</td>
</tr>
<tr>
<td>Cloud cover (0–4hal)</td>
<td>2.8 (2.0–3.7)</td>
<td>2.4 (1.5–3.0)</td>
<td>–</td>
</tr>
<tr>
<td>Air temperature, °C</td>
<td>21.6 (19.2–23.9)</td>
<td>25.2 (23.3–26.6)</td>
<td>21.5 (19.0–24.0)</td>
</tr>
<tr>
<td>Water temperature, °C</td>
<td>19.0 (17.6–20.9)</td>
<td>21.3 (20.3–22.5)</td>
<td>18.0 (16.0–20.0)</td>
</tr>
<tr>
<td>Wind speed, m/s</td>
<td>7.0 (5.3–8.2)</td>
<td>6.6 (5.3–7.3)</td>
<td>4.9 (2.7–5.5)</td>
</tr>
</tbody>
</table>

**Season**, n (%)

<table>
<thead>
<tr>
<th></th>
<th>High (187 (41.1))</th>
<th>Low (268 (58.9))</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>90 (83.3)</td>
<td>314 (61.2)</td>
</tr>
<tr>
<td></td>
<td>199 (38.8)</td>
<td>25 (24.0)</td>
</tr>
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</table>

**Type of day, n (%)**

<table>
<thead>
<tr>
<th></th>
<th>Weekday (213 (46.8))</th>
<th>Weekend (82 (18.0))</th>
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</thead>
<tbody>
<tr>
<td></td>
<td>18 (16.7)</td>
<td>233 (45.4)</td>
</tr>
<tr>
<td></td>
<td>233 (45.4)</td>
<td>88 (17.2)</td>
</tr>
<tr>
<td></td>
<td>12 (11.1)</td>
<td>17 (16.3)</td>
</tr>
</tbody>
</table>

**Vacation, n (%)**

<table>
<thead>
<tr>
<th></th>
<th>78 (22.2)</th>
<th>192 (37.4)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>80 (23.1)</td>
<td>63 (60.6)</td>
</tr>
</tbody>
</table>

Wave factor: wave height (m) times wave period (s).
The wave incidence factor ranges from 0 to 1; see equation (1).
Forecast values not shown because of differences in the modes of data measurement.
Significant differences between observed and forecast data.
High season: 15 June to 15 September.

The retrospective and prospective cohort studies included 325 (246) rescues and 95 (71) drownings on 106 (84) days. The 3-hour model was well calibrated in terms of retrospective data (goodness-of-fit test p=0.20 for the 3-hour model, p=0.53 for the 3-hour-step model). Both models exhibited significant p values on Spiegelhalter Z-testing of prospective data, evidencing a lack of calibration: the daily model tended to overpredict days with risks of drowning >0.5; the 3-hour step model overpredicted risks as low as 0.1.

Using prospective data with 3-day forecasts, we found that assessment of the coast-wide risk of water adverse event using the five-level scale missed 1 of 158 days featuring a rescue at
the lowest risk level (0.6%). The missed case was a rescued man who presented without a cough and was discharged on site. The prospective data predicted 45.8% of days with rescue events at the highest risk level (table 3). Few differences between forecasts according to their delay and the risk level were observed (see online supplemental material table A2). The 3-hour step model missed 2 of 481 rescues, one at the lowest level (0.4%) and one at the highest (15.7%).

Observed rescues and drownings by predicted risk level derived using regression models exploiting 3-day forecasts; Gironde, southwestern France.

The missed case at the lowest risk level in the retrospective cohort occurred during moderate wave conditions (wave factor $\sim 8 \text{m} \times \text{s}$) and at low wave incidence factor (0.47), but the victim required only rescue, was asymptomatic on rescue, and was not evacuated. The second missed event occurred at the tip of the Cap Ferret sandspit, which lacks wave-driven rip currents. The last missed event occurred at La Salie Nord, adjacent to (south of) the Arcachon inlet, under moderate wave conditions.

DISCUSSION

While previous studies examined the association of ocean drowning risk with weather conditions, our study is the first to focus on the prediction of life-risk. Air temperature, wave

<table>
<thead>
<tr>
<th>Predictor</th>
<th>Crude OR (95% CI)</th>
<th>Adj. OR (95% CI)</th>
<th>$\chi^2$</th>
<th>dof</th>
</tr>
</thead>
<tbody>
<tr>
<td>Wave factor*, m×s</td>
<td></td>
<td></td>
<td></td>
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<tr>
<td>&lt;5.2</td>
<td>Ref.</td>
<td>Ref.</td>
<td></td>
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<tr>
<td>5.2–9.2</td>
<td>3.89 (2.03 to 9.81)</td>
<td>6.41 (2.93 to 18.3)</td>
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<tr>
<td>&gt;9.2</td>
<td>1.96 (1.01 to 4.50)</td>
<td>7.10 (3.09 to 22.7)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Wave incidence factor*</td>
<td>1.83 (1.40 to 2.62)</td>
<td>2.27 (1.58 to 3.70)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Cloud cover†</td>
<td>0.52 (0.37 to 0.74)</td>
<td>–</td>
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<tr>
<td>Air temperature*, °C</td>
<td></td>
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<tr>
<td>≤21</td>
<td>Ref.</td>
<td>Ref.</td>
<td></td>
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<tr>
<td>21–23.5</td>
<td>6.58 (3.03 to 19.2)</td>
<td>4.79 (1.93 to 16.6)</td>
<td></td>
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</tr>
<tr>
<td>&gt;23.5</td>
<td>19.10 (9.62 to 61.8)</td>
<td>12.20 (4.69 to 52.3)</td>
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<tr>
<td>Water temperature*, °C</td>
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<tr>
<td>≤19.5</td>
<td>Ref.</td>
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<tr>
<td>19.5–21.3</td>
<td>3.95 (2.13 to 9.14)</td>
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<tr>
<td>&gt;21.3</td>
<td>13.18 (7.60 to 27.0)</td>
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<td></td>
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<tr>
<td>Wind speed, m/s</td>
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<td></td>
<td></td>
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<tr>
<td>≤4.3</td>
<td>Ref.</td>
<td>–</td>
<td></td>
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<tr>
<td>4.3–6.3</td>
<td>1.71 (1.00 to 3.14)</td>
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<td></td>
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<tr>
<td>&gt;6.3</td>
<td>0.84 (0.45 to 1.66)</td>
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<tr>
<td>Season</td>
<td></td>
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<tr>
<td>Low</td>
<td>Ref.</td>
<td>Ref.</td>
<td></td>
<td></td>
</tr>
<tr>
<td>High</td>
<td>7.17 (4.38 to 13.5)</td>
<td>3.98 (1.44 to 6.11)</td>
<td></td>
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<tr>
<td>Type of day</td>
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<tr>
<td>Weekday</td>
<td>Ref.</td>
<td>Ref.</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Weekend</td>
<td>1.73 (0.71 to 3.62)</td>
<td>2.96 (1.06 to 7.98)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Holidays</td>
<td>5.77 (3.49 to 10.8)</td>
<td>4.25 (2.19 to 9.75)</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

Univariate and multivariate analyses performed with the aid of logistic regression models using retrospective data from 2011 to 2013.

*Daily maximal value.
†Daily mean value.
‡Not incorporated into multivariate analyses because of differences in measurement modes.
direction and the wave factor were the primary environmental predictors; the type of day and the season were also significant, but less important, predictors. Given the availability of extensive metocean data used to infer rip current activity, we were able to build a tool that accurately predicted coast-wide life risk.

Warm weather increases sea exposure and therefore the risk of drowning, consistent with other studies. This finding, associated with the importance of the time of day, highlights the social dimension of the drowning risk. Actually, air temperature and time of the day may be the primary controls for the number of beach users and their behaviours.

Wave parameters influencing rip current flow velocity were significant predictors of drowning, consistent with the results of physical models. For instance, coast-wide life-risk was increased for more shore-normally incident waves. Here, the life-risk was found to systematically increase with the wave factor $W_f$, which is in line with Castelle et al., but contrast with Scott et al. in UK where hazard was maximised for just below average $W_f$. In a previous study, it was qualitatively shown that the beach morphology, which has a profound impact on rip current hazard, was important to the number of drowning incidents. Given the low number of studied seasons and events, and the likely contrasting beachgoer profiles between seasons, it was not possible to robustly address the influence of the bathymetry. A beach bathymetry proxy could be added to the model via the integration of beach state estimation model based on the equilibrium concept.

As the wind parameters differed between the retrospective and prospective periods, we could not use these parameters, although they might have further improved the models. Cloud cover was measured differently during the two periods and thus could not be incorporated into the models. Although univariate analysis showed that cloud cover was a significant predictor of drowning, it is strongly correlated with air temperature. Future models should integrate predicted rather than observed measures. Parameters such as cloud cover could be therefore studied.

Our models tend to overestimate the risk on days associated with moderate to high risks; some variables may thus be unknown, related to the beach morphology and the beach attendance. First, the summer beach morphology along the Gironde coast is very variable with rip channel exhibiting different morphologies from one summer to another, and even in space with more gently sloping and less channelized beaches northwards. Given that, for instance, the relative depth of the rip channels is critical to rip current flow speed, and thus physical hazard, and that beach type also influence attendance, consideration of a constant beach morphology is limiting. Although the current version of our model does not account for beach attendance data, it can guide lifeguard/rescue decision makers who need to allocate resources. Moreover, drownings are certainly under-reported to the emergency call centre; reporting rates may vary over time. We could not directly estimate exposure, as beach attendance is not measured in Gironde. This findings might also be due to the consistency between the perception of the hazard by beach users and the computed risk. The timestamps for the calls are provided automatically by the emergency call centre database. It may be delayed by a few minutes from the event.

Turning to the missed events, two occurred in sectors adjacent to the Arcachon lagoon inlet, where local, strong tide-driven currents develop at low tide, constituting a major hazard. We hypothesise that the missed events were attributable to these currents. This highlights the need to carefully target preventative messages; the primary hazards vary locally.

Use of the 1, 2 and 3-day forecasts yielded similar results; this will aid in the efficient deployment of lifeguards and rescue equipment. Accurate local forecasts more than 3 days ahead are not available.

How may our findings save lives? This work can be used in two ways: the first as a trigger for targeted preventative messages, and the second as a decision aid for rescue services.

The use of a binary scale would trigger many false alarms; thus, we considered that a five-level scale was more appropriate, as such scales are used to predict other risks posed by natural hazards (such as snow avalanches). Our scale should be improved using a risk utility function, which remains to be specifically determined. We concede that our present levels are arbitrary; we must still explore what beachgoers and decision-makers consider to be ‘low’, ‘moderate’ or ‘acceptable’ risks. The study of social factors influencing beach attendance and water use might still be important to study to formulate preventative messages. This will be the object of a future study.

Any message suggested by our models must be consistent with ‘messages’ imparted by beach flags. These flags can be ‘green’ (no or minor hazard, bathing supervised), ‘yellow’ (hazard, bathing supervised) or ‘red’ (major hazard, no bathing allowed). They are determined by lifeguards based on wave conditions, water temperature, and beach attendance, and may vary depending (for example) on lifeguard experience. Unpublished local reports indicate that green and red flags are rarely raised during the high season on the Gironde coast. The future integration of forecasts into drowning prevention strategies should therefore take into account the social dimension of rip current hazard.

We are confident that our model can be adapted to similar beaches with rip currents, but complete generalisation of our findings is inappropriate in the absence of more data. Lifeguard knowledge and the physical parameters of natural hazards require attention: our model was built on lifeguard hypothesis for beach attendance and physics models of natural hazard. The next steps are forecast validation by lifeguards and automation of feedback; these will allow the model to be continuously improved. As more data become available, other modelling strategies may be appropriate, such as Bayesian or neural networks. In such strategies, the importance of the exposure would be of interest to identify environmental conditions prone to risky behaviour.

Predicting the need for rescue from water in a hazardous environment is key to reducing the risk of drowning. Our predictive models can be used to efficiently deploy medical teams and

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### Table 3: Observed rescues versus predicted coast-wide risk of adverse water event

<table>
<thead>
<tr>
<th>Risk level</th>
<th>Events*</th>
<th>Total</th>
<th>%</th>
<th>Events</th>
<th>Total</th>
<th>%</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>1</td>
<td>112</td>
<td>2.7</td>
<td>1</td>
<td>172</td>
<td>0.6</td>
</tr>
<tr>
<td>2</td>
<td>1</td>
<td>112</td>
<td>0.9</td>
<td>8</td>
<td>105</td>
<td>7.1</td>
</tr>
<tr>
<td>3</td>
<td>11</td>
<td>113</td>
<td>9.7</td>
<td>15</td>
<td>95</td>
<td>15.8</td>
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<td>4</td>
<td>3</td>
<td>112</td>
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<td>23</td>
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<td>5</td>
<td>70</td>
<td>113</td>
<td>61.9</td>
<td>66</td>
<td>144</td>
<td>45.8</td>
</tr>
</tbody>
</table>

*Events: rescues and drownings.
†Goodness of fit: $p=0.53$, calibration test $p=0.10$. Brier scores: 0.05 for 2011–2013, 0.05 for 2015–2017.
‡Goodness of fit: $p=0.20$, calibration test $p<0.001$. Brier scores: 0.10 for 2011–2013, 0.12 for 2015–2017.
What is already known on the subject

- Drownings along surf beaches are mainly caused by rip currents.
- Rip currents activity is controlled by wave and beach morphology factors.

What this study adds

- Along surf beaches of south western France, drownings can be anticipated because of wave and wave forecast in combination with calendar factors.
- Such methods could be adapted to other places with similar hazards.

rescue helicopters. An interventional study (performed under real-world conditions) is planned. A utility function reflecting risk perception/acceptance is required. This would allow prioritised preventative messages to be broadcast during high-risk periods. The strategy must employ behavioural change theory to reduce the risk to beachgoers.

Evaluation requires reliable data from both lifeguard stations and emergency call centre files.

Contributors ÉT: conceptualisation, methodology, software, validation, formal analysis, investigation, data curation, original draft preparation, review and editing. ML-B: methodology, original draft preparation, review. BC: conceptualisation, data curation, methodology, original draft preparation, review. BS: conceptualisation, methodology, data curation, investigation, review. RS: original draft preparation, review and editing.

Funding The authors have not declared a specific grant for this research from any funding agency in the public, commercial or not-for-profit sectors.

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Competing interests None declared.

PATIENT CONSENT FOR PUBLICATION Not required.

Ethics approval Data collection was approved by the French national committee protecting data privacy (Commission Nationale de l’Informatique et des Libertés, CNIL), provided that only compiled (anonymised) data would be published. French law states that a retrospective observational study does not require ethics committee approval.

Provenance and peer review Not commissioned; externally peer reviewed.

Data availability statement Compiled data are available in a public, open access repository. Detailed data may be obtained from a third party and are not publicly available. Therefore, daily model data are available on the OSF repository: https://osf.io/4bQa/ and, due to data providers restrictions, 3-hour step model data are available upon request from the corresponding author (eric.tellier@chu-bordeaux.fr).

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