Traffic calming policy can reduce inequalities in child pedestrian injuries: database study

S J Jones, R A Lyons, A John, S R Palmer

Objectives: To determine whether area wide traffic calming distribution reflects known inequalities in child pedestrian injury rates. To determine whether traffic calming is associated with changes in childhood pedestrian injury rates.

Design: Small area ecological study, longitudinal analysis of injury rates with cross sectional analysis of traffic calming and method of travel to school.

Settings: Two cities in the United Kingdom.


Main outcome measures: Area wide traffic calming distribution by area deprivation status and changes in injury rate/1000.

Results: The most deprived fourth of city A had 4.8 times (95% CI 3.71 to 6.22) the number of traffic calming features per 1000 population compared with the most affluent fourth. Injury rates among the most deprived dropped from 9.42 to 5.07 from 1992–94 to 1998–2000 (95% CI for change 2.82 to 5.91). In city B, the traffic calming ratio of the most to least deprived fourth was 1.88 (95% CI 1.46 to 2.42); injury rates in the deprived areas dropped from 8.92 to 7.46 (95% CI for change −0.84 to 3.77). Similar proportions of 9–12 year olds walked to school in both cities.

Conclusions: Area wide traffic calming is associated with absolute reductions in child pedestrian injury rates and reductions in relative inequalities in child pedestrian injury rates.
Traffic calming audit

A street by street audit of traffic calming was carried out during April and May 2002 in both cities. Traffic calming features audited were speed humps, road narrowings, and road closures. Twelve figure grid references were obtained using global positioning systems (GPS) and plotted using ArcView 3.2.

Initially it was intended that traffic calming feature installation dates would be used in a before and after approach. However, local authorities do not store these data in a readily available manner. Also, traffic calming has been installed on different roads in the same EDiv over a period of months or years, precluding simple time series analyses. Installation was mainly from the mid 1990s onwards; we therefore used three year periods from 1992 to 2000.

Pedestrian casualties

Injurious road traffic accidents, recorded by the police STATS19 form, were accessed through the UK data archive. Child pedestrian casualty locations in both cities for the entire study period were mapped to EDivs and deprivation fourths. Three year data aggregations were used to generate sufficient cases to improve precision of estimates.

Population denominators

The 4–16 year old population of each EDiv and fourth was calculated using an extract from the NHS Administrative Register. To account for population changes, 1991 census data were aggregated to 1998 EDiv boundaries and population denominators for 1992 (denominator for 1992–94) and 1996 (denominator for 1995–97) estimated. For 1998–2000 the 1999 NHSAR population estimate was the denominator.

Analysis of traffic calming and injuries data

Traffic calming “rates” were calculated using the 1999 4–16 year population as a denominator and number of calming features as numerator. Traffic calming “rate” ratios were calculated for each fourth against the most affluent.

Child pedestrian injury rates were calculated using the STATS19 numerator and population denominators. Changes in injury rates between 1992–94 and 1998–2000 were calculated, with 95% confidence intervals. Injury rate ratios, against the most affluent, were calculated for each fourth.

Potential confounders

To account for possible confounding data on school travel mode was sought. Few appropriate exposure data were already available, but the road safety department of city A was independently conducting a school travel survey. We conducted a similar survey in city B. A single sheet survey was used, with pupils aged 9–12 years asked about modes of travel on that day. The proportions of children travelling to school by each mode and deprivation fourth were calculated. Car ownership data were obtained from the 1991 and 2001 censuses.

The total lengths of all roads, by fourth, were calculated using Meridian 2 software. Total road length, average distance between features, and feature density (features per unit road length) were calculated.

RESULTS

Traffic calming

There were 891 traffic calming features in city A, mainly in the most deprived fourth (71.3%; table 1). City B had 553 features, with 48.3% in the most deprived fourth. There were 4.8 times more features in the most deprived compared with most affluent fourth of City A (table 1).

Road network

Features in city A were closer than city B for all except the most affluent (table 2). Features were eight times further apart in the most affluent areas, compared with the most deprived in city A. In city B the ratio was 4.2.

<table>
<thead>
<tr>
<th>Table 2</th>
<th>Road network measures (km) by deprivation fourth (source: Meridian 2.0, ONS, 2003)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>City A</td>
</tr>
<tr>
<td></td>
<td>Overall average distance between features (km)</td>
</tr>
<tr>
<td>Most affluent</td>
<td>240.3</td>
</tr>
<tr>
<td>Next affluent</td>
<td>300.9</td>
</tr>
<tr>
<td>Most deprived</td>
<td>114.6</td>
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<tr>
<td>Total</td>
<td>733.1</td>
</tr>
</tbody>
</table>
Injury rates

For 1992–94, child pedestrian injury rates were 6.98 per 1000 4–16 year olds in city A and 6.05 in city B (table 3). By 1998–2000, city A’s injury rate had dropped significantly (decrease 2.14), while a non-significant change of 0.8 occurred in city B.

Most deprived injury rates also dropped significantly in city A (table 3); the inequalities gap between the most affluent and most deprived reduced from 3.2 in 1992–94 to 2.0 in 1998–2000. In city B, there were no significant injury rate changes and little change in inequalities.

Changes in injury rates were significantly inversely correlated with density of traffic calming features (number of features per km road length; \( r = 0.769, p = 0.026 \); table 2).

School travel

In city A, 61.5% of most deprived children walked to school and 64.2% in city B.

Table 3 Injuries, injury rates (per 1000 4–16 year olds, per three years), and changes in injury rates (1998–2000 v 1992–94) by deprivation fourth

<table>
<thead>
<tr>
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</tr>
</thead>
<tbody>
<tr>
<td>Most affluent</td>
<td>n = 37</td>
<td>n = 33</td>
<td>n = 41</td>
<td>-0.11 (-1.46 to 1.19)</td>
<td>n = 15</td>
<td>n = 15</td>
<td>n = 15</td>
<td>-0.17 (-1.74 to 1.38)</td>
</tr>
<tr>
<td>Next affluent</td>
<td>n = 17</td>
<td>n = 17</td>
<td>n = 9</td>
<td>-2.42 (-5.29 to -0.10)</td>
<td>n = 25</td>
<td>n = 25</td>
<td>n = 25</td>
<td>-1.48 (-3.70 to 0.70)</td>
</tr>
<tr>
<td>Next deprived</td>
<td>n = 74</td>
<td>n = 74</td>
<td>n = 62</td>
<td>-0.94 (-2.18 to 0.26)</td>
<td>n = 73</td>
<td>n = 73</td>
<td>n = 73</td>
<td>-3.74 (-6.20 to 2.74)</td>
</tr>
<tr>
<td>Most deprived</td>
<td>n = 217</td>
<td>n = 175</td>
<td>n = 151</td>
<td>-3.68 (-5.28 to -2.13)</td>
<td>n = 111</td>
<td>n = 96</td>
<td>n = 94</td>
<td>-1.32 (-3.62 to 0.97)</td>
</tr>
<tr>
<td>Total</td>
<td>n = 345</td>
<td>n = 305</td>
<td>n = 268</td>
<td>-2.14 (-2.81 to -1.48)</td>
<td>n = 224</td>
<td>n = 217</td>
<td>n = 201</td>
<td>-0.80 (-1.88 to 0.27)</td>
</tr>
</tbody>
</table>

Rate ratio most deprived v most affluent

| City A | 3.21 (2.27 to 4.54) |
| City B | 2.92 (2.02 to 4.24) |

Note: CI, confidence interval.

DISCUSSION

The distribution of traffic calming in two UK cities favours disadvantaged children, although to a much greater extent in city A. If traffic calming reduces pedestrian injuries (the cited literature suggests some doubt), then areas with more calming should see greater reductions in injuries, all else being equal. Our results are compatible with this explanation with greater reductions in deprived areas of city A, where there was more traffic calming, and a significant correlation between change in pedestrian injury rates and density of traffic calming. The phrase ‘all else being equal’ is particularly significant as injury rate changes could be due to measured or unmeasured confounders.

This type of n = 2 ecological study suffers from lack of power and the potential for bias due to other simultaneous changes in intervention or control areas. However, this study design is almost inevitable in policy evaluation because of difficulties with getting different jurisdictions to randomly allocate policies and laws. Evaluations of many interventions face these difficulties—for example, mandatory use of helmets. Therefore, authors need to consider all possible biases, the extent to which they have or can be measured, and the plausible magnitude and effect of any residual biases on results before accepting that change is due to the intervention.

The main weaknesses of this study were the lack of traffic calming installation dates. Neither city council held a comprehensive listing of installed features, their locations, and installation dates. Each feature had a design plan, held in hundreds of separate files. In both cities, installation dates were available from the mid 1990s onwards. Even if installation dates were available, it would not have helped evaluation because features were installed on different roads in individual EDivs over periods of months or years. Evaluating the impact of individual features on child pedestrians is difficult because the number of injuries adjacent to any one structure is often 0 or 1; our 1998–2000 injury rate was 0.3 injuries per kilometre.

Other potential biases relate to changing population and environmental exposures between cities, and between deprived fourths within cities. Population denominators were specific to each time period. Injury rates therefore accounted for population changes.

For 1992–94, child pedestrian injury rates were 6.98 per 1000 4–16 year olds in city A and 6.05 in city B (table 3). By 1998–2000, city A’s injury rate had dropped significantly (decrease 2.14), while a non-significant change of 0.8 occurred in city B.

Most deprived injury rates also dropped significantly in city A; the inequalities gap between the most affluent and most deprived reduced from 3.2 in 1992–94 to 2.0 in 1998–2000. In city B, there were no significant injury rate changes and little change in inequalities.

Changes in injury rates were significantly inversely correlated with density of traffic calming features (number of features per km road length; \( r = 0.769, p = 0.026 \); table 2).
Total road length did not differ substantially between cities. Ideally, these calculations would count only roads suitable for traffic calming and along which child pedestrians are exposed, but such decisions are subjective.

Traffic volumes and speeds may have changed during the study period, but there are no suitable data to assess this possibility. National estimates are available but are not appropriate for regional comparisons. Some route analyses have been carried out within each city, but these focus on specific roads for a short period of time—periodic repeat data collection is very unusual.

Exposure to traffic could also have changed or may have varied considerably between the two cities. A National Travel Survey comparing 1989–91 with 1998–2000 found that walking to school by primary school children dropped from 62% to 56%, and from 48% to 43% for secondary school children. These changes are not sufficient to explain the injury rate changes in city A and should have lead to similar injury rate decreases in both cities. Also, similar proportions of children walked to school in both cities; higher injury rates in city B are not due to greater exposure. It is possible that there were substantial differences at the beginning of the study period, but discussions with the road safety departments in both cities suggested that no obvious changes had occurred.

When this study began, city A was collecting exposure data for “safe routes to schools” projects which were being implemented in all cities from 2000 onwards; we used an identical form in city B. Data were collected for 9–12 year olds and the sole purpose of collecting these data was to determine whether differences in child pedestrian exposures may explain some of these findings. Analysis of only the 9–12 year age group was considered, but the number of cases (n = 421) would have considerably reduced the precision of the point estimates, making definitive conclusions even more difficult. Furthermore, the most appropriate analysis could be argued to include only school journey injuries (n = 265). Generalising these data to all 4–16 year olds and all journeys is not ideal and limits these analyses. However, collecting even these data was a substantial task. The alternative would have been to ignore possible variations in exposures and be left with even more unanswered questions. Few road traffic environment studies have adequate exposure data to address all potential confounding.

The limitations of STATS19 have been widely discussed. \(^22\)–\(^24\) The main concerns include case ascertainment completeness and accuracy of grid reference reporting. Child pedestrian casualties are underreported in police statistics. \(^24\)–\(^27\) For this reason, and because of the inaccuracies in police statistics, the main concern includes case ascertainment completeness and accuracy of grid reference reporting. Child pedestrian casualties are underreported in police statistics. \(^24\)–\(^27\) The limitations of STATS19 have been widely discussed. \(^22\)–\(^24\)

Definitions of casualties as fatal, serious, or slight are also known to be problematic. \(^24\) For this reason, and because of the small numbers likely in the fatal and serious category, no severity analysis was carried out. More discrepancies are likely when comparing data from different police forces, but both cities are in the same police force area.

Police officers do not carry GPS readers, making inaccuracies in accident location recording possible. In addition, STATS19 is often completed some time after the incident and only records a 10 figure grid reference. Therefore the potential exists for error in accident location recording, but this would be of greater concern with an individual feature based before and after approach. With the fourths ecological approach, systematic errors within and between cities are less likely.

STATS19 only records accident location, leading to questions over population denominators. The majority of child pedestrians are injured within 1 km of their home and the residential population of the area where the accident occurred is therefore an appropriate denominator. \(^27\)–\(^30\) There may be some age related variations in these distances, with older children travelling further than younger children. However, the issue here is not the actual distance travelled, but whether the child crosses an EDiv boundary during the journey, thereby making the residential population an inappropriate denominator. Child pedestrian injuries have previously been shown to cluster around the home and occur in the census tract where the child lives, regardless of age. \(^29\)–\(^30\)

These findings have been presented to the engineering and road safety departments of both cities. Discussions also took place regarding traffic calming installation and child road safety policies, such as safe routes to school. Some differences in the mechanism for prioritising traffic calming sites were identified, but there were no obvious differences in road safety education that could have changed behaviors and reduced injury risk in city A. \(^22\)–\(^27\)

A final issue is whether regression to the mean could explain these changes. This occurs when the group being tested has high values initially, increasing the likelihood that subsequent values will be lower. Since initial injury rates in both cities were similar, it is unlikely that regression towards the mean could explain the differential changes. Regression to the mean may account for changes within each city, changes that we are attributing to traffic calming in city A. However, this argument requires that regression to the mean takes place over a period of nine years. The length of an appropriate period for a regression to the mean consideration is open to debate, as is the point in time at which the mean is calculated.

Drawing definitive conclusions from these analyses is difficult. The ecological approach means that unmeasured factors could have accounted for injury rate changes. However, given the observed changes in inequalities and the analysis of the potential confounders presented here, substantial influence of other variables seems unlikely. Previous research has studied the effectiveness of traffic calming, but none has tried to examine the effects on pedestrian injury inequalities.

Other than a large multisite randomised controlled trial, it is difficult to imagine an evaluation framework which could produce unequivocal results. In the absence of sufficient financial and political support for the large scale and expensive trials there is merit in replicating this study in other areas.

ACKNOWLEDGEMENTS

Thanks to road safety engineers in both cities for general advice and to Sylvie Morgan, Linda Jones, Brian Jones, and Luke Jones for assistance with surveying.

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This research was partly funded by a grant from the Chief Medical Officer for Wales.
First step toward mandatory US standard for cigarette lighters

The US Consumer Product Safety Commission (CPSC) has voted unanimously to move forward with the first of three steps in developing a new mandatory safety standard for cigarette lighters. The vote to approve an Advanced Notice of Proposed Rulemaking sets the Commission on a path to consider a way to prevent most mechanical malfunctions of lighters and reduce the fire hazard associated with some lighters. CPSC already has a mandatory standard for child-resistant cigarette lighters which addresses the hazard of child deaths under 5 years starting fires with lighters. That standard for child-resistance applies to imported as well as domestically-manufactured disposable and novelty lighters. There are nearly one billion cigarette lighters sold in the US annually. Over 700 million lighters are imported each year, with about 400 million coming from China. From 1997 to 2002, the CPSC estimated that more than 3000 people went to hospital emergency rooms for injuries resulting from malfunctioning lighters. Most of these injuries involved thermal burns to the face, hands, and fingers. The voluntary standard for lighters addresses the risk of fire, death, and injury associated with mechanical malfunctions of lighters. However, it is unclear how many lighters sold in the US actually comply with the voluntary standard. Fire deaths associated with children playing with lighters dropped dramatically since the mandatory standard for child-resistance became effective in July 1994—from 230 in 1994 to 130 in 1998. Children under 5 years accounted for 170 of the deaths in 1994 and 40 of the deaths in 1998. In 1994 there were 11 100 residential fires associated with children playing with lighters. By 1998, that number declined to 6100 fires.

REFERENCES